

# On the Frequency Distribution of Non-Independent Random Values

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## Abstract

Let  $\mathbf{Z} = (Z_1, \dots, Z_n)$  be an  $n$ -tuple of random variables and let  $\mathcal{P}$  be a convex set containing all conditional probability distributions  $P_{Z_i|Z_1=z_1 \dots Z_{i-1}=z_{i-1}}$ . We show that the frequency distribution of the elements in the  $n$ -tuple  $\mathbf{Z}$  is contained in an  $\varepsilon$ -environment of  $\mathcal{P}$ , except with small probability.

## 1 Preliminaries

**Definition 1.1.** The *variational distance* between two probability distributions  $P$  and  $Q$  over the same range  $\mathcal{Z}$  is defined by

$$\delta(P, Q) := \frac{1}{2} \sum_{z \in \mathcal{Z}} |P(z) - Q(z)|.$$

**Definition 1.2.** Let  $\varepsilon \geq 0$ . The  $\varepsilon$ -environment  $\mathcal{B}_\varepsilon(\mathcal{P})$  of a set of probability distributions  $\mathcal{P}$  with range  $\mathcal{Z}$  is the set of probability distributions  $P'$  with range  $\mathcal{Z}$  such that  $\delta(P, P') \leq \varepsilon$  for some  $P \in \mathcal{P}$ .

**Definition 1.3.** Let  $\mathbf{z} := (z_1, \dots, z_n)$  be an  $n$ -tuple of elements of a set  $\mathcal{Z}$ . The *frequency*  $Q_{\mathbf{z}}$  of  $\mathbf{z}$  is the function from  $\mathcal{Z}$  to  $[0, 1]$  defined by

$$Q_{\mathbf{z}}(z) := \frac{|\{i : z_i = z\}|}{n},$$

for  $z \in \mathcal{Z}$ .

It is easy to see that the frequency  $Q_{\mathbf{z}}$  is a probability distribution on  $\mathcal{Z}$ .

**Definition 1.4.** A *martingale* is an  $n + 1$ -tuple  $(Z_0, \dots, Z_n)$  of random variables on  $\mathbb{R}$  such that,

$$E[Z_i | Z_1 = z_1, \dots, Z_{i-1} = z_{i-1}] = z_{i-1}$$

for any  $i \in \{1, \dots, n\}$  and  $z_1, \dots, z_{n-1} \in \mathcal{Z}$  (if  $\text{Prob}[Z_1 = z_1, \dots, Z_{i-1} = z_{i-1}] > 0$ ).

**Theorem 1.5 (Azuma's Inequality).** Let  $(Z_0, \dots, Z_n)$  be a martingale with  $Z_0 = 0$  and

$$|Z_i - Z_{i-1}| \leq 1$$

for all  $i \in \{1, \dots, n\}$ . Then, for any  $\mu > 0$ ,

$$\text{Prob}[Z_n > \mu\sqrt{n}] < e^{-\mu^2/2} .$$

*Proof.* See, e.g., p. 95 of [1]. □

## 2 Main Theorem and Proof

**Theorem 2.1.** Let  $\mathbf{Z} = (Z_1, \dots, Z_n)$  be an  $n$ -tuple of random variables with alphabet  $\mathcal{Z}$  and let  $\mathcal{P}$  be a convex set of probability distributions on  $\mathcal{Z}$  such that

$$P_{Z_i|Z_1=z_1, \dots, Z_{i-1}=z_{i-1}} \in \mathcal{P}$$

for all  $i \in \{1, \dots, n\}$  and  $z_1, \dots, z_{i-1} \in \mathcal{Z}$  (if  $\text{Prob}[Z_1 = z_1, \dots, Z_{i-1} = z_{i-1}] > 0$ ). Then, for any  $\varepsilon \geq 0$ ,

$$\text{Prob}[Q_{\mathbf{Z}} \in \mathcal{B}_\varepsilon(\mathcal{P})] > 1 - 2^{|\mathcal{Z}|} e^{-n\varepsilon^2/2} .$$

*Proof.* Let  $\mathcal{A}$  be any fixed nonempty proper subset of  $\mathcal{Z}$ . For  $i \in \{1, \dots, n\}$ , let  $B_i$  be the binary random variable which takes the value 1 if  $Z_i \in \mathcal{A}$  and 0 otherwise, let  $p_i$  be the function on  $\mathcal{Z}^{i-1}$  defined by

$$p_i(z_1, \dots, z_{i-1}) := \text{Prob}[B_i = 1 | Z_1 = z_1, \dots, Z_{i-1} = z_{i-1}] ,$$

and let

$$D_i := B_i - p_i(Z_1, \dots, Z_{i-1}) .$$

For  $i \in \{0, \dots, n\}$ , let

$$S_i := \sum_{j=1}^i D_j .$$

We show that the  $n+1$ -tuple  $(S_0, \dots, S_n)$  is a martingale, i.e.,

$$E[S_i | S_0 = s_0, \dots, S_{i-1} = s_{i-1}] = s_{i-1} \tag{1}$$

for all  $i \in \{1, \dots, n\}$  and  $s_0, \dots, s_{n-1}$ . Since  $B_i$  is binary, we have

$$E[B_i | Z_1 = z_1, \dots, Z_{i-1} = z_{i-1}] = p_i(z_1, \dots, z_{i-1}) ,$$

and, by the definition of  $D_i$ ,

$$E[D_i | Z_1 = z_1, \dots, Z_{i-1} = z_{i-1}] = 0 .$$

The values of  $(D_1, \dots, D_n)$  are uniquely determined by the values of  $(Z_1, \dots, Z_n)$ . We thus have

$$E[D_i | D_1 = d_1, \dots, D_{i-1} = d_{i-1}] = 0$$

and, using the definition of  $S_i$ ,

$$E[S_i | D_1 = d_1, \dots, D_{i-1} = d_{i-1}] = E\left[\sum_{j=1}^i D_j | D_1 = d_1, \dots, D_{i-1} = d_{i-1}\right] = \sum_{j=1}^{i-1} d_j .$$

which implies (1). Since  $S_0 = 0$  and  $|S_i - S_{i-1}| = |D_i| \leq 1$  for all  $i \in \{1, \dots, n\}$ , we can apply Theorem 1.5 leading to

$$\text{Prob}[S_n > \mu\sqrt{n}] < e^{-\mu^2/2} . \quad (2)$$

For any  $\mathcal{A} \subseteq \mathcal{Z}$ , let  $s_{\mathcal{A}}$  be the function

$$s_{\mathcal{A}} : P \mapsto \sum_{z \in \mathcal{A}} P(z) ,$$

defined on the set of probability distributions on  $\mathcal{Z}$ . For any  $\mathbf{z} = (z_1, \dots, z_n)$ , let  $P_{\mathbf{z}}$  be the probability distribution on  $\mathcal{Z}$  defined by

$$P_{\mathbf{z}}(z) := \frac{1}{n} \sum_{j=1}^n \text{Prob}[Z_j = z | Z_1 = z_1, \dots, Z_{j-1} = z_{j-1}] \quad (3)$$

for  $z \in \mathcal{Z}$ . Since, by the definition of  $\mathbf{B} = (B_1, \dots, B_n)$ ,

$$s_{\mathcal{A}}(Q_{\mathbf{z}}) = Q_{\mathbf{B}}(1) = \frac{1}{n} \sum_{j=1}^n B_j$$

and, by the definition of  $p_j$  and  $P_{\mathbf{z}}$ ,

$$s_{\mathcal{A}}(P_{\mathbf{z}}) = \frac{1}{n} \sum_{j=1}^n p_j(Z_1, \dots, Z_{j-1})$$

we find

$$s_{\mathcal{A}}(Q_{\mathbf{z}}) - s_{\mathcal{A}}(P_{\mathbf{z}}) = \frac{1}{n} \sum_{j=1}^n (B_j - p_j(Z_1, \dots, Z_{j-1})) = \frac{1}{n} \sum_{j=1}^n D_j = \frac{1}{n} S_n .$$

It follows from (2) that

$$\text{Prob}[s_{\mathcal{A}}(Q_{\mathbf{z}}) - s_{\mathcal{A}}(P_{\mathbf{z}}) > \frac{\mu}{\sqrt{n}}] < e^{-\mu^2/2} . \quad (4)$$

The variational distance between  $Q_{\mathbf{z}}$  and  $P_{\mathbf{z}}$  can be written as

$$\delta(Q_{\mathbf{z}}, P_{\mathbf{z}}) = \max_{\mathcal{A} \subseteq \mathcal{Z}} (s_{\mathcal{A}}(Q_{\mathbf{z}}) - s_{\mathcal{A}}(P_{\mathbf{z}}))$$

where the maximum ranges over all non-empty proper subsets  $\mathcal{A}$  of  $\mathcal{Z}$ . Applying the union bound for all  $2^{|\mathcal{Z}|} - 2$  subsets  $\mathcal{A}$ , we conclude from (4)

$$\text{Prob}[\delta(Q_{\mathbf{z}}, P_{\mathbf{z}}) > \frac{\mu}{\sqrt{n}}] < (2^{|\mathcal{Z}|} - 2)e^{-\mu^2/2} < 2^{|\mathcal{Z}|} e^{-\mu^2/2} .$$

The assertion then follows from the observation that the probability distribution  $P_{\mathbf{z}}$ , as defined in (3), is a convex combination of distributions from the set  $\mathcal{P}$ , i.e., since  $\mathcal{P}$  is convex,  $P_{\mathbf{z}} \in \mathcal{P}$ , for any  $\mathbf{z}$ , and, consequently,  $P_{\mathbf{z}} \in \mathcal{P}$ .  $\square$

## References

- [1] N. Alon and J. H. Spencer. *The probabilistic method*. Wiley-Interscience Series in Discrete Mathematics and Optimization. Wiley, second edition, 2000.